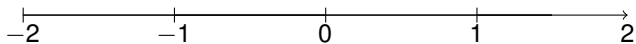


# Large deviations for a random walk in random environment revisited

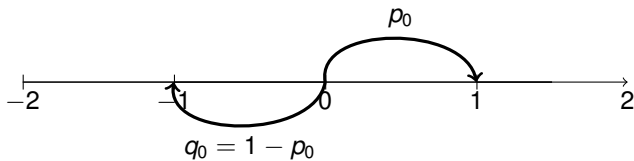
Piotr Dyszewski  
University of Wrocław  
(joint with Dariusz Buraczewski)

Będlewo, 7.05.2018

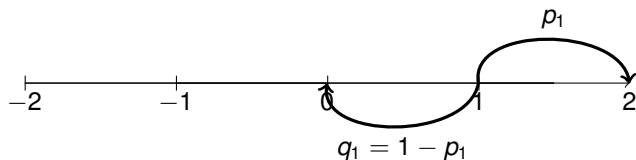
$\omega = \{p_k\}_{k \in \mathbb{Z}}$ ,  $p_k \in (0, 1)$  are iid.



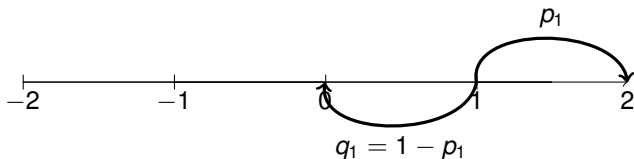
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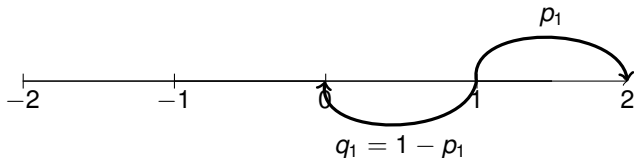


$X = \{X_n\}_{n \in \mathbb{N}}$  is a random walk in random environment (RWRE)

$$P_\omega[X_{n+1} = k + 1 | X_n = k] = p_k$$

$$P_\omega[X_{n+1} = k - 1 | X_n = k] = q_k.$$

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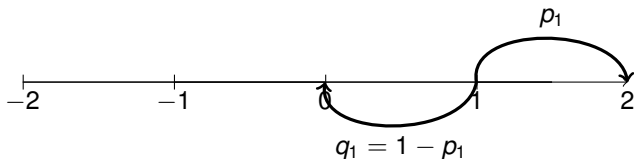
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$P_\omega$  – quenched probability.

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$P_\omega$  – quenched probability. Define the annealed probability  $\mathbb{P}$  viz.

$$\mathbb{P}[X \in A, \omega \in B] = \int_B P_\omega[X \in A] P(d\omega).$$

Let  $A = \frac{q}{p}$ . Assume

$$\mathbb{E}A^\alpha = 1 \quad \alpha > 1$$

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Solomon (Ann. Probab. '75)

$$\lim_{n \rightarrow \infty} \frac{X_n}{n} = v = \frac{1 - \mathbb{E}A}{1 + \mathbb{E}A}$$

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$$\lim_{n \rightarrow \infty} \frac{X_n}{n} = v = \frac{1 - \mathbb{E}A}{1 + \mathbb{E}A} \neq \mathbb{E}p - q$$

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Kesten, Kozlov, Spitzer (Compos. Math. '75)

$$\alpha \in (1, 2), \quad \frac{X_n - vn}{n^{\frac{1}{\alpha}}} \Rightarrow \mathcal{L}_\alpha, \quad \alpha > 2, \quad \frac{X_n - vn}{n^{\frac{1}{2}}} \Rightarrow \mathcal{N}$$

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Dembo, Peres, Zeitouni (Commun. Math. Phys. '96)

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Buraczewski, D. (submitted '17)

$$\mathbb{P}[X_n < (v - \varepsilon)n] \sim C\varepsilon^{-\alpha} n^{1-\alpha}$$

$\mathbb{E}A^\alpha = 1$  for  $\alpha > 1$ . Let  $T_n = \inf\{k : X_k = n\}$ .

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$$\frac{X_n}{n} \sim v \iff \frac{T_n}{n} \sim \frac{1}{v}$$

$$\mathbb{P}\left[\frac{X_n}{n} < v - \varepsilon\right] \sim C_\varepsilon n^{1-\alpha} \iff \mathbb{P}\left[\frac{T_n}{n} > \frac{1}{v} + \varepsilon\right] \sim C'_\varepsilon n^{1-\alpha}$$

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$$\mathbb{P}\left[\frac{X_n}{n} < \nu - \varepsilon\right] \sim C_\varepsilon n^{1-\alpha} \iff \mathbb{P}\left[\frac{T_n}{n} > \frac{1}{\nu} + \varepsilon\right] \sim C'_\varepsilon n^{1-\alpha}$$

$T_n = \#$  of steps during  $[0, T_n)$

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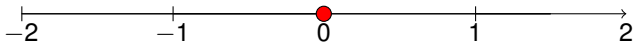
$T_n = \#$  of steps during  $[0, T_n)$   
=  $\#$  of steps to the right during  $[0, T_n)$   
+  $\#$  of steps to the left during  $[0, T_n)$

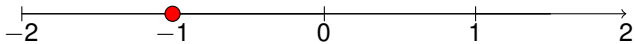
$\mathbb{E}A^\alpha = 1$  for  $\alpha > 1$ . Let  $T_n = \inf\{k : X_k = n\}$ . Then

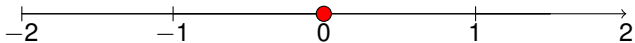
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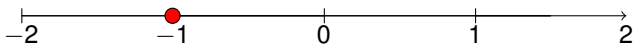
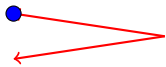
$$\mathbb{P}\left[\frac{X_n}{n} < v - \varepsilon\right] \sim C_\varepsilon n^{1-\alpha} \iff \mathbb{P}\left[\frac{T_n}{n} > \frac{1}{v} + \varepsilon\right] \sim C'_\varepsilon n^{1-\alpha}$$

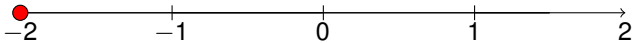
$$\begin{aligned} T_n &= \# \text{ of steps during } [0, T_n) \\ &= \# \text{ of steps to the right during } [0, T_n) \\ &\quad + \# \text{ of steps to the left during } [0, T_n) \\ &= n + 2 \cdot \# \text{ of steps to the left during } [0, T_n) \end{aligned}$$

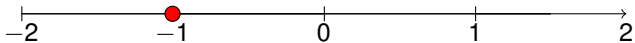
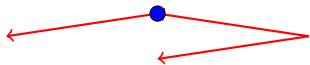


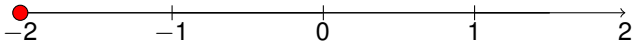


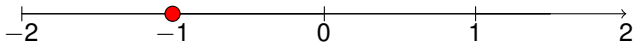
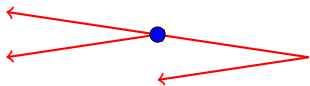


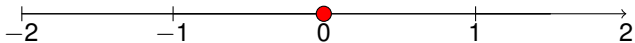


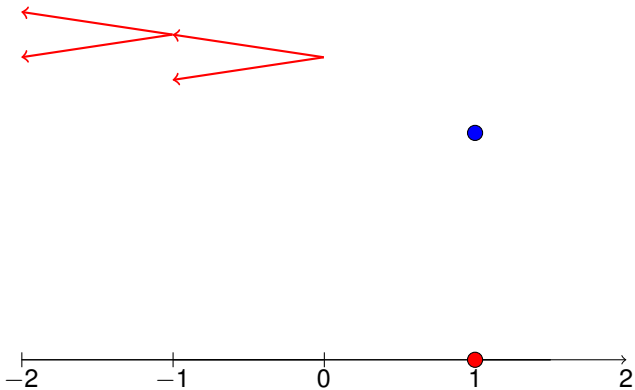


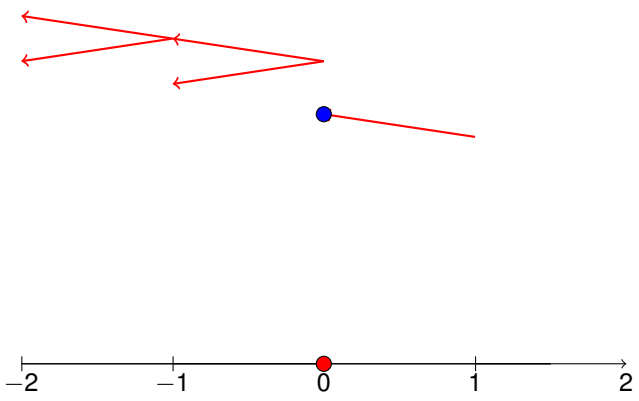


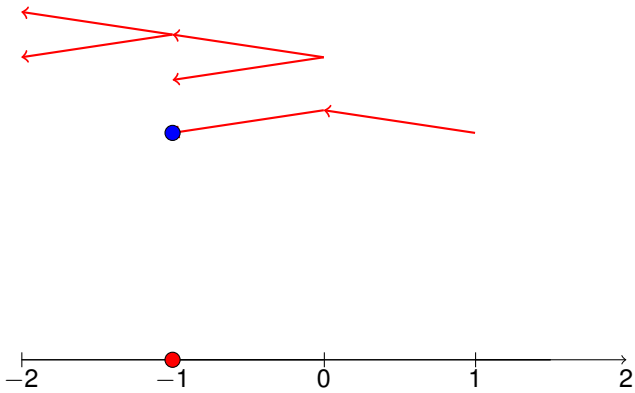


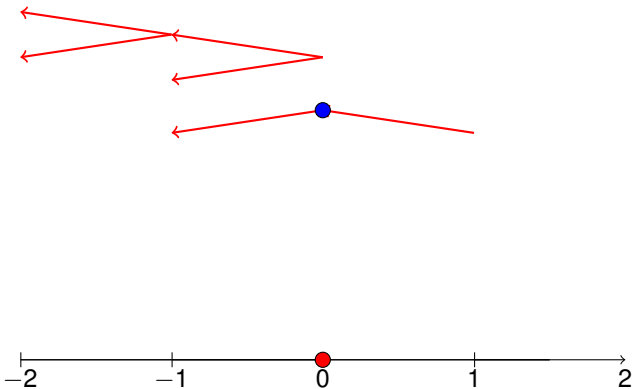


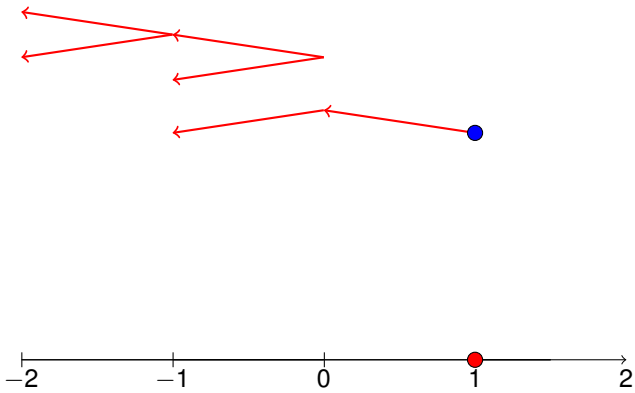


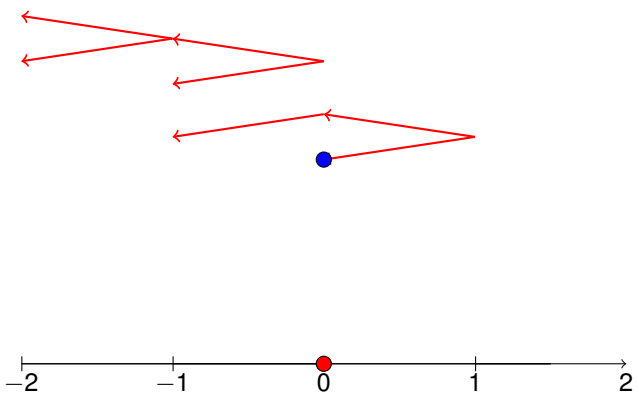


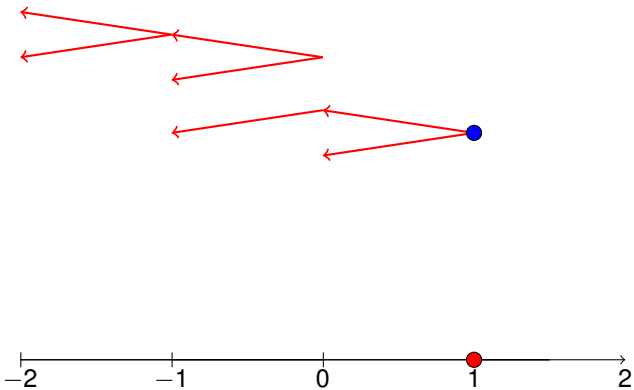


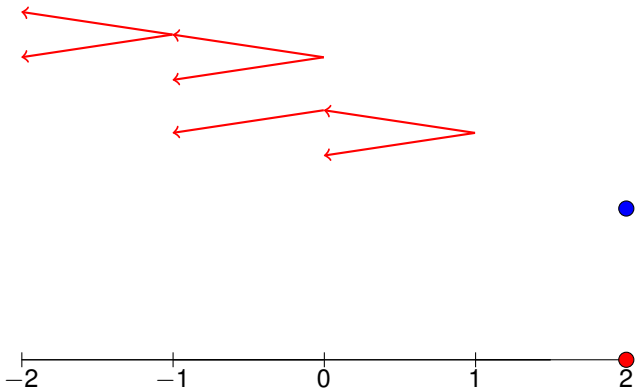


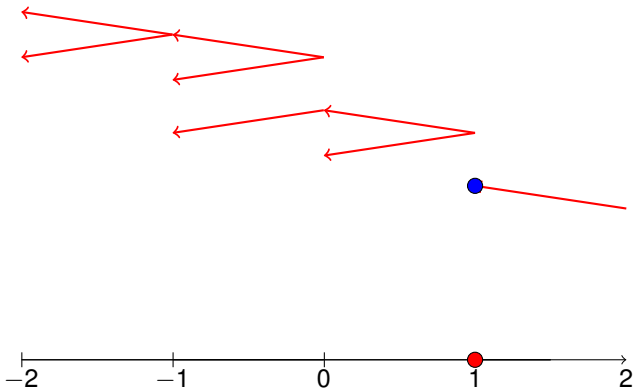


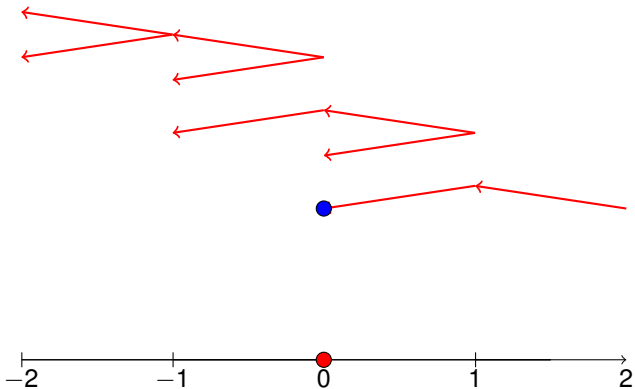


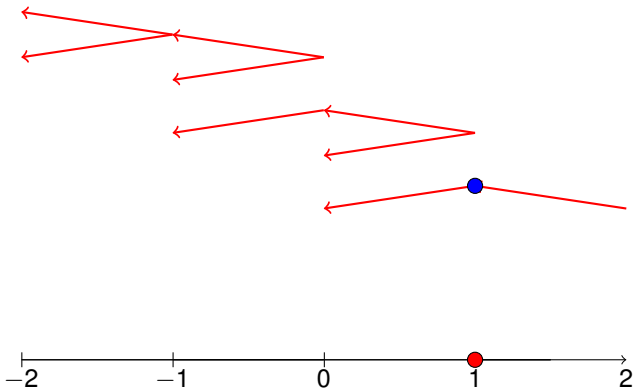


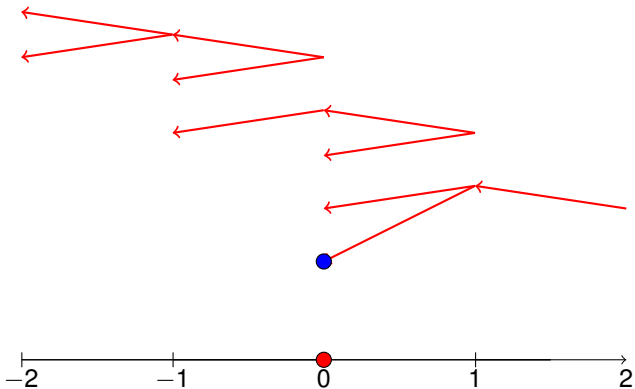


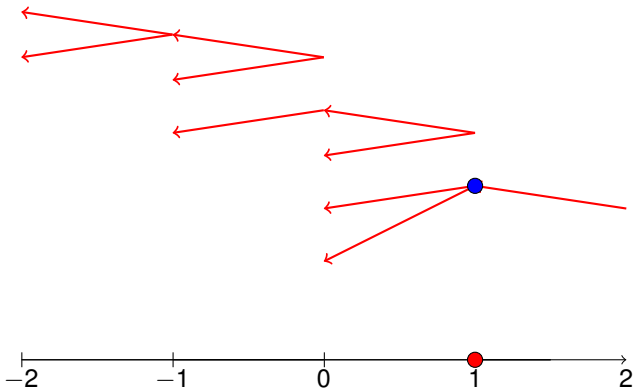


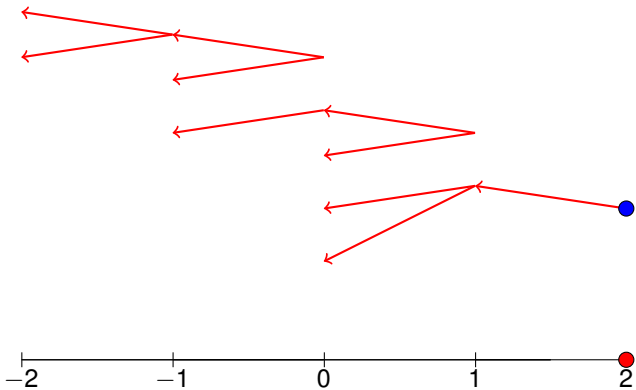


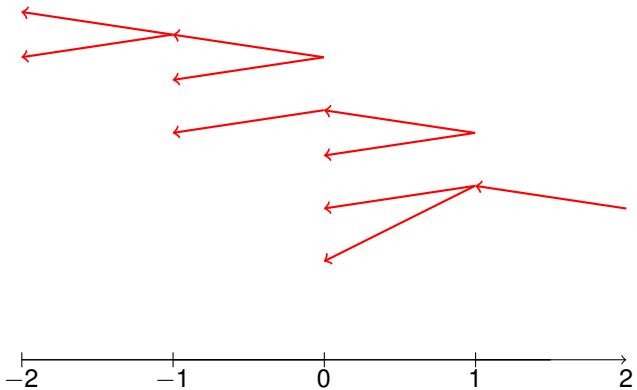


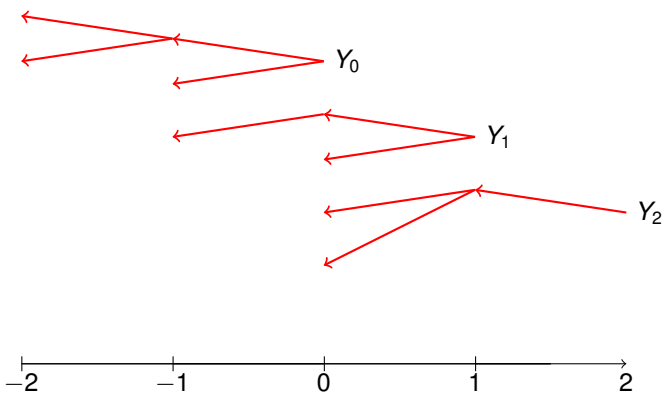


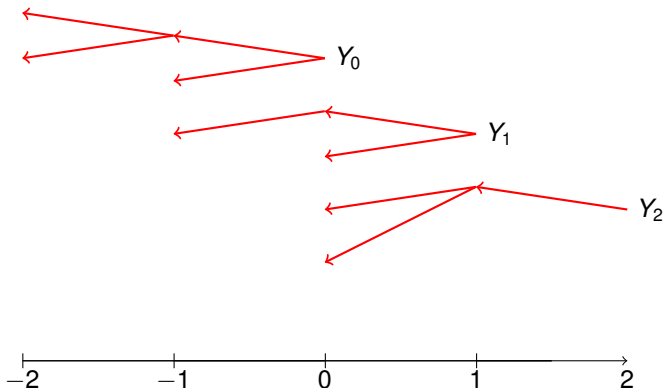




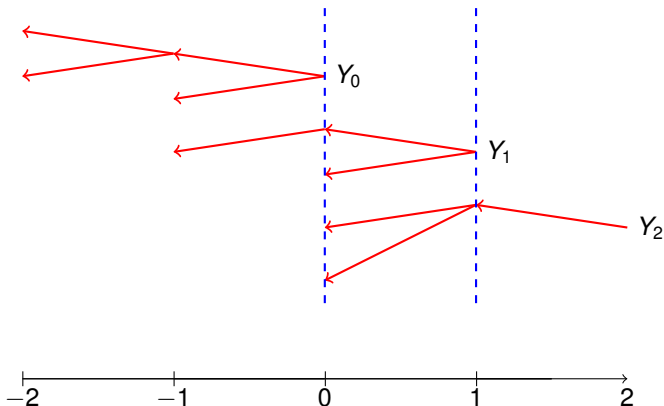




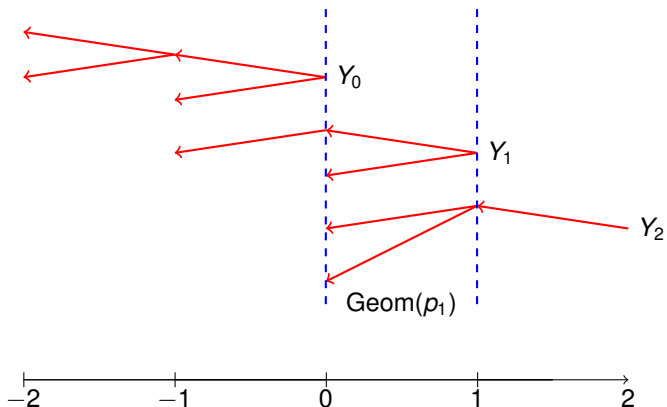




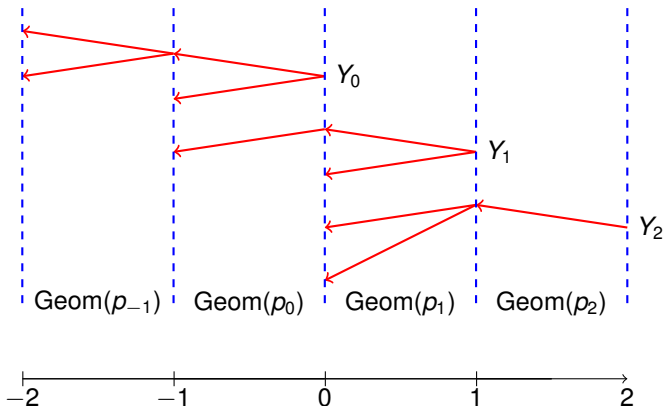
$$\# \text{ of steps to the left during } [0, T_n) = \sum_{k=0}^{n-1} Y_k$$



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**subcritical BPRE**  $Y_0 = \sum_{k=1}^{\infty} Z_{0,k}$

$$Z_{0,0} = 1$$

$$Z_{0,k} = \sum_{j=1}^{Z_{0,k-1}+1} V_j^k,$$

where  $V_j^k \sim \text{Geom}(p_k)$ ,  $P_{\omega}[V_j^k = m] = q^m p$ ,  $\mathbb{E}_{\omega} V_j^k = \frac{q_k}{p_k} = A_k$

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Kesten, Kozlov, Spitzer (Compos. Math. '75)

$$\mathbb{P}[Y_0 = \sum_{k=0}^{\infty} Z_{0,k} > cn] \sim C_Y n^{-\alpha}$$

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Buraczewski, D. (submitted '17)  $\rho = \mathbb{E}A^{\alpha} \log A$ ,

$$C_Y n^{-\alpha} \sim \mathbb{P}\left[\sum_{k=0}^{\infty} Z_{0,k} > cn\right] \sim \mathbb{P}\left[\sum_{k=n_1}^{n_2} Z_{0,k} > cn\right],$$

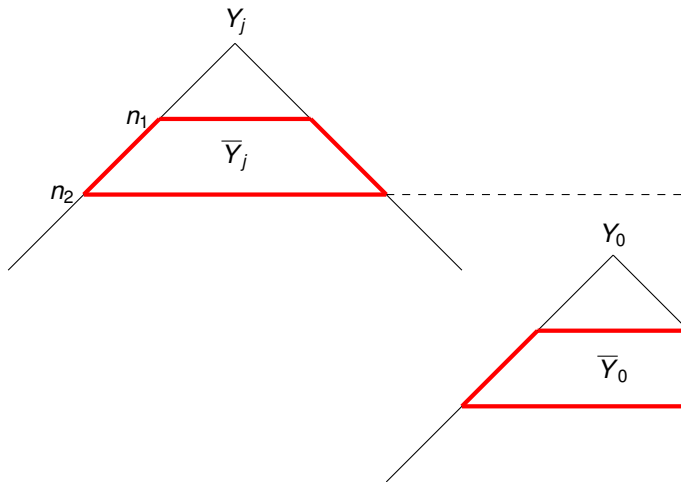
and

$$\mathbb{P}\left[\sum_{k \notin [n_1, n_2]} Z_{0,k} > cn\right] = o(n^{-\alpha})$$

where  $n_1 = \frac{1}{\rho} \log n - \sqrt{\log n}$ ,  $n_2 = \frac{1}{\rho} \log n + \sqrt{\log n}$ .

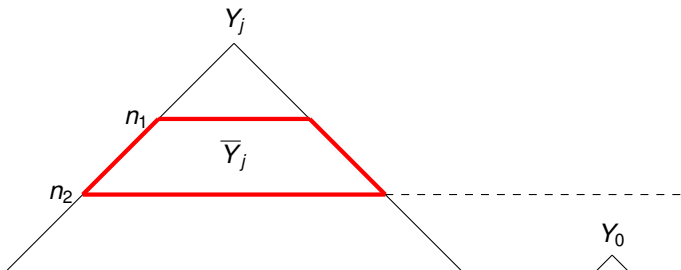
$$Y_0 \approx \sum_{n_1 \leq k < n_2} Z_{0,k} = \bar{Y}_0$$

$$Y_j \approx \sum_{n_1+j \leq k < n_2+j} Z_{j,k} = \bar{Y}_j$$

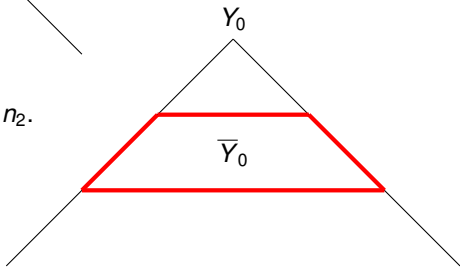


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$\bar{Y}_0$  and  $\bar{Y}_j$  are independent if  $j > n_2$ .



$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right]$$

$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right]$$

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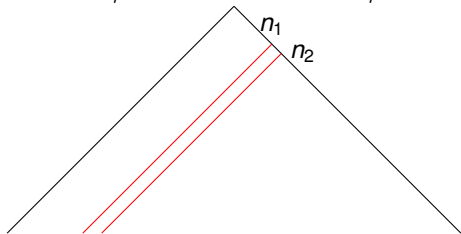
for  $n_1 = \frac{1}{\rho} \log n - \sqrt{\log n}$ ,

$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right]$$

for  $n_1 = \frac{1}{\rho} \log n - \sqrt{\log n}$ ,  $n_2 = \frac{1}{\rho} \log n + \sqrt{\log n}$ .

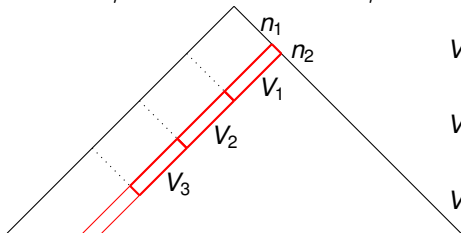
$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right]$$

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$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right]$$

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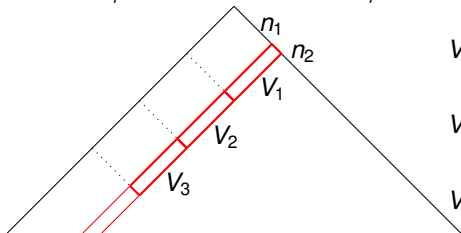
$$V_1 = \sum_{0 \leq i < n_2} \bar{Y}_i$$

$$V_2 = \sum_{n_2 \leq i < 2n_2} \bar{Y}_i$$

$$V_3 = \sum_{2n_2 \leq i < 3n_2} \bar{Y}_i$$

$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right]$$

for  $n_1 = \frac{1}{\rho} \log n - \sqrt{\log n}$ ,  $n_2 = \frac{1}{\rho} \log n + \sqrt{\log n}$ .



$$V_1 = \sum_{0 \leq i < n_2} \bar{Y}_i$$

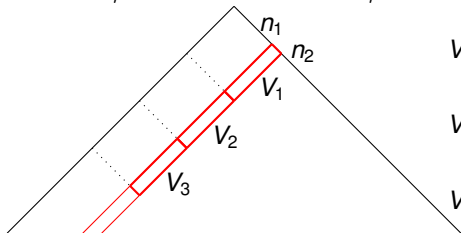
$$V_2 = \sum_{n_2 \leq i < 2n_2} \bar{Y}_i$$

$$V_3 = \sum_{2n_2 \leq i < 3n_2} \bar{Y}_i$$

$V_1$  depends on  $p_0, p_1, \dots, p_{2n_2-1}$ ,  $V_3$  depends on  $p_{2n_2}, p_{2n_2+1}, \dots, p_{4n_2-1}$ .

$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right]$$

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**$V_1$  and  $V_3$  are independent and identically distributed!**

Buraczewski, Damek, Mikosch, Zienkiewicz (Ann. Probab. '13)  
Precise large deviations for 1-dependent sequence

$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{0 \leq i \leq n} \bar{Y}_i > cn\right] = \mathbb{P}[V_1 + \dots + V_{\frac{n}{n_2}} > cn]$$

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$$\mathbb{P}\left[\sum_{i=0}^{n-1} Y_i > cn\right] \sim \mathbb{P}\left[\sum_{i=1}^{n/n_2} V_i > cn\right] \sim \frac{n}{n_2} \mathbb{P}[V_1 > cn] \sim Cn^{1-\alpha}$$

Let  $A = \frac{q}{p}$ . Assume

$$\mathbb{E}A^\alpha = 1 \quad \alpha > 1$$

Solomon (Ann. Probab. '75)

$$\lim_{n \rightarrow \infty} \frac{X_n}{n} = v = \frac{1 - \mathbb{E}A}{1 + \mathbb{E}A} > 0$$

Kesten, Kozlov, Spitzer (Compos. Math. '75)

$$\alpha \in (1, 2), \quad \frac{X_n - vn}{n^{\frac{1}{\alpha}}} \Rightarrow \mathcal{L}_\alpha, \quad \alpha > 2, \quad \frac{X_n - vn}{n^{\frac{1}{2}}} \Rightarrow \mathcal{N}$$

Dembo, Peres, Zeitouni (Commun. Math. Phys. '96)

$$\log \mathbb{P}[X_n < (v - \varepsilon)n] \sim (1 - \alpha) \log n.$$

Buraczewski, D. (submitted '17)

$$\mathbb{P}[X_n < (v - \varepsilon)n] \sim C\varepsilon^{-\alpha} n^{1-\alpha}$$



D. Buraczewski, E. Damek, T. Mikosch, and J. Zienkiewicz.

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D. Buraczewski and P. D.

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H. Kesten, M. V. Kozlov, and F. Spitzer.

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F. Solomon.

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$$C_Y n^{-\alpha} \sim \mathbb{P}\left[\sum_{k=0}^{\infty} Z_{0,k} > cn\right] \sim \mathbb{P}\left[\sum_{k=n_1}^{n_2} Z_{0,k} > cn\right],$$

where  $n_1 = \frac{1}{\rho} \log n - \sqrt{\log n}$ ,  $n_2 = \frac{1}{\rho} \log n + \sqrt{\log n}$ .

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Buraczewski, D. (submitted '17)

$$\frac{\tau_n}{\log(n)} \Big|_{\tau_n < \infty} \rightarrow \mathbb{P} \frac{1}{\rho}.$$



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